

A Rapid Acquisition Technique for Impulse Radio

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Abstract—A novel rapid acquisition algorithm, called *Sequential Block Search*, is proposed for impulse radio systems, which can reduce the mean acquisition time considerably. The algorithm is compared to the conventional serial search algorithm and output statistics for both algorithms are derived for a binary phase-shift keyed random time hopping impulse radio system. The mean acquisition time formulas are obtained using the signal flow graph approach. The simulations indicate the superiority of the rapid acquisition algorithm over the serial search technique with a modest increase in the complexity.

I. INTRODUCTION

In impulse radio communication systems, the acquisition of an incoming signal is critical in de-modulation of the signal which often takes the form of PPM (pulse position modulation) or BPSK (binary phase shift keying). Because of the high resolution in time required to locate narrow pulses employed in ultra-wideband (UWB) systems, a considerable number of possible pulse positions must be searched in order to acquire the received signal. Therefore, a rapid acquisition algorithm is very important in UWB communications.

The role of an acquisition technique is essentially to determine the relative delay of the received signal with respect to a template signal in the receiver. The conventional technique to achieve this is what is known as the serial search algorithm [1], in which the received signal is correlated with the template signal and the correlation output is compared to a threshold. If the output is lower than the threshold, the template signal is shifted by an amount comparable to pulse interval and is correlated with the received signal. This procedure continues until an output exceeds the threshold. This serial search algorithm may take a long time, thus, some quick algorithm is called for.

We propose a rapid acquisition algorithm, called *Sequential Block Search* (SBS), which can considerably reduce the mean acquisition time, while maintaining the receiver complexity comparable to the serial search algorithm.

The SBS algorithm first aims to determine a subregion where the signal delay is likely to exist. Then, this subregion is examined in a finer detail to find the exact position of the signal. In order to perform the first step quickly, we define a new template signal. The correlation of the received signal with this template signal effectively adds a number of serial search

correlation outputs within a certain block in time. This output from the block is then used as a criterion to decide if the block contains the signal or not. Whenever the block output exceeds a specified threshold, we apply the conventional serial search algorithm within this subregion corresponding to the block.

The remainder of the paper is organized as follows. Section II defines the signal structure used during the acquisition process. SBS algorithm is explained in Section III. Then, statistics of correlation outputs are derived in Section IV, which is followed by acquisition time analysis and simulation results in Section V. Finally, some concluding remarks are made in Section VI.

II. SIGNAL MODEL

Consider a binary phase-shift keyed random time hopping impulse radio (TH-IR) system where the transmitted signal from user k in an N_u -user setting is represented by the following model [3]:

$$s_{tr}^k(t) = \sum_{j=-\infty}^{\infty} d_j^k b_{[j/N_f]}^k w_{tx}(t - jT_f - c_j^k T_c), \quad (1)$$

where w_{tx} is the transmitted unit-energy pulse, T_f is the average pulse repetition time, N_f is the number of pulses representing one information symbol and $b_{[j/N_f]}^k \in \{+1, -1\}$ is the information symbol transmitted by user k . In order to allow the channel to be exploited by many users and avoid catastrophic collisions, a pseudo-random sequence $\{c_j^k\}$, where $c_j^k \in \{0, 1, \dots, N_c - 1\}$, is assigned to each user. This sequence is called time hopping sequence and provides an additional time shift of $c_j^k T_c$ seconds to the j th pulse of the k th user where T_c is the chip interval and is chosen to satisfy $T_c \leq T_f/N_c$ in order to prevent the pulses from overlapping.

In this paper, we consider coded IR systems where d_j^k 's are binary random variables, independent for $(j, k) \neq (i, l)$, taking each of the values ± 1 with probability $1/2$ [3]. This system can be regarded as a random Direct Sequence CDMA (RCDMA) system with $T_f = T_c$ and N_f representing the processing gain.

Define a sequence $\{s_j^k\}$ as follows

$$s_j^k = \begin{cases} d_{[j/N_c]}^k, & j - N_f [j/N_c] = c_{[j/N_c]}^k \\ 0, & \text{otherwise.} \end{cases} \quad (2)$$

Then, assuming $T_f/T_c = N_c$ without loss of generality, (1)

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can be expressed as

$$s_{tr}^k(t) = \sum_{j=-\infty}^{\infty} s_j^k b_{\lfloor j/(N_f N_c) \rfloor} w_{tx}(t - jT_c), \quad (3)$$

which indicates that IR systems can be regarded as RCDMA systems with generalized spreading sequences, $\{s_j^k\}$ [3].

Assume that no data modulation is done during the acquisition stage, that is, $b_{\lfloor j/(N_f N_c) \rfloor} = 1 \forall j$ ³. In this case, the received signal over a flat fading channel⁴ can be expressed as

$$r(t) = \sum_{k=1}^{N_u} \sum_{j=-\infty}^{\infty} s_j^k w_{rx}(t - jT_c - \tau_k) + \sigma_n n(t), \quad (4)$$

where τ_k is the TOA of user k ($\tau_k = lT_c$, with l being an integer, is assumed for simplicity), w_{rx} is the received UWB pulse and $n(t)$ is white Gaussian noise with unit spectral density.

III. SEQUENTIAL BLOCK SEARCH ALGORITHM

First, consider the conventional serial search algorithm. In this scheme, the received signal is correlated with a template signal and the output is compared to a threshold. If the output is lower than the threshold, the template signal is shifted by one bin, which corresponds to the resolvable path interval and the correlation with the received signal is obtained again. By this way, the search continues until an output exceeds the threshold.

Each different search location is called a "cell". If the output of the correlation comes from a case where signal path and the template signal are aligned, it is called a signal cell output. Otherwise, it is called a non-signal cell output. The number of cells in the uncertainty region is taken to be N . One of these cells is the signal cell while the others are non-signal cells. A false alarm occurs when a non-signal cell output exceeds the threshold. In this case, time t_p elapses until the search recovers again. This time is called penalty time for a false alarm.

For the signal model in (3), the template signal for the serial search can be expressed as follows, assuming no data modulation for the purposes of acquisition:

$$s_{m_1}^{(c)}(t) = \sum_{n=jN_c}^{(j+m_1)N_c-1} s_n^1 w_{rx}(t - nT_c), \quad (5)$$

where m_1 is the number of pulses, over which the correlation is taken.

For the SBS algorithm, there are two different template signals. The first one is similar to the one used in the serial search while the second one is used for searching a block of cells quickly. This latter template signal for the signal model described in (3) can be expressed as follows:

$$s_{m_2}^{(b)}(t) = \sum_{i=0}^{K-1} \sum_{n=jN_c}^{(j+m_2)N_c-1} s_n^1 w_{rx}(t - nT_c - iT_c), \quad (6)$$

³Without loss of generality $b_{\lfloor j/(N_f N_c) \rfloor} = 1 \forall j$, k can be assumed for our purposes.

⁴The results for frequency selective channels can be found in [2].

where K is the number of cells in the block and $N = KB$ is assumed with B being the total number of blocks in the uncertainty region. For simplicity, T_c is taken as the minimum resolvable path interval.

The output of the correlation between the received signal and the template signal in (6) is used as a quick test to check if the whole block consisting of K cells contains the signal cell or not, while the correlation output of the received signal and the first template in (5) is used in the detailed search of a block.

Let i denote the index of the block that is currently being searched, with $i = 1$ initially. Then, the SBS algorithm can be described as follows:

- 1) Check the i th block using $s_{m_2}^{(b)}(t - (i-1)KT_c)$.
- 2) If the output of the i th block is not higher than the threshold, τ_b , then go to step 6.
- 3) If the output of the i th block is higher than τ_b , then search the block in more detail, i.e. cell-by-cell serial search with threshold τ_s and template $s_{m_1}^{(c)}(t)$.
- 4) If no signal cell is found in the block, go to step 6.
- 5) If the signal cell is found in the block, DONE.
- 6) Set $i = (i \bmod B) + 1$ and go to step 1.

Some implicit steps in the algorithm are worth considering. When a false alarm (FA) occurs in the serial search part, the search resumes with the next cell after C time units, which is the penalty time in terms of frame time. So in step 5, "the signal cell is found" means that a correlation output exceeds the threshold without a false alarm. Similarly, in step 4, "no signal cell is found" implies that no correlation output exceeds the threshold without a false alarm.

IV. STATISTICS OF THE DECISION VARIABLES

In this section, the statistics of the decision variables will be derived for long correlation intervals.

A. Serial Search Outputs

In this case, the received signal is correlated with shifted versions of the template signal (5).

$$y_i = \int r(t) s_{m_1}^{(c)}(t - \Delta_i) dt, \quad (7)$$

for $i = 1, \dots, N$, where $\Delta_i = (i-1)T_c$ is the delay of the template signal for the i th cell. The output y_i can be expressed as the sum of three terms, $y_i = x_i + a_i + n_i$, where

- x_i : Signal part/Self interference,
- a_i : Multiple access interference (MAI),
- n_i : Output noise.

The output noise can be shown to be distributed as $\mathcal{N}(0, m_1 \sigma_n^2)$. The first term x_i can be expressed as follows depending on the signal/non-signal cell situation:

1) *Signal Cell Case*: In this case, the template signal (5) and the received signal are aligned. Therefore, the correlation over m_1 pulses gives the following signal part

$$x_i = \sum_{j=1}^{m_1 N_c} s_j^2 = m_1, \quad (8)$$

since s_j is non-zero (± 1) only once in N_c consecutive cells. Note that l_s is the index of the signal cell.

2) *Non-Signal Cell Case*: Again consider a correlation over m_1 pulses. In this case, the received signal and the template signal in (5) are misaligned by one or more cells. Then, the self interference part of the output of the correlation can be expressed as

$$x_i = \sum_{n=jN_c}^{(j+m_1)N_c-1} \int s_n^1 w_{rx}(t-nT_c-\Delta_i) \sum_{m=-\infty}^{\infty} s_m^1 w_{rx}(t-mT_c-\tau_1) dt \quad (9)$$

for $i = 1, \dots, l_s - 1, l_s + 1, \dots, N$.

Note that since the outer summation is over $m_1 N_c$ terms, s_n^1 will be non-zero (± 1) just m_1 times. So (9) can be considered as the sum of m_1 separate correlations, each of which denotes the output due to one pulse of the template signal. If we denote each of these terms by a random variable x_{il} for $l = 1, \dots, m_1$, the self interference part of the output can be expressed as

$$x_i = \sum_{l=1}^{m_1} x_{il}, \quad (10)$$

where

$$x_{il} = \int d_{j+l-1}^1 w_{rx}(t-nT_c-\Delta_i) \sum_{m=-\infty}^{\infty} s_m^1 w_{rx}(t-mT_c-\tau_1) dt. \quad (11)$$

Now the aim is to find the probability distribution of x_{il} for each l . To this end, assume that the amount of misalignment between the received signal and the template signal is an integer multiple of T_c for all non-signal cell situations, that is, $\tau_1 - \Delta_i$ is an integer multiple of T_c . Let $\gamma_i = |\tau_1 - \Delta_i|/T_c$ be the integer representing the misalignment between the delay of the i th template signal and the received signal in terms of T_c , which can also be interpreted as the distance between the i th non-signal cell and the signal cell.

Note that since each x_{il} is the output of the correlation between one pulse in the template signal and the received signal, it is obvious that x_{i1}, \dots, x_{im_1} are identically distributed.

First, consider the case where $\gamma_i < N_c$, that is, the misalignment between the received and the template signals is less than $N_c T_c$ seconds. In this case, x_{il} is non-zero only when the l th pulse of the template signal overlaps with any of the received signal pulses. Since $\gamma_i < N_c$, two pulses can overlap only if they are in frames which are next to each other. Therefore, for the l th pulse of the template signal to overlap with a pulse of the received signal, the pulse should be in one of the γ_i positions close to the next frame and the next pulse in that frame should be in the position, to which this l th pulse is shifted. The probability that these two events occur simultaneously is γ_i/N_c^2 . Therefore, with probability γ_i/N_c^2 , the output is non-zero. This non-zero output is the multiplication of two consecutive coding values: d_{j+l-1}^1 and d_{j+l}^1 or d_{j+l-1}^1 and d_{j+l-2}^1 , where d_j^k is as in (1) and related to the spreading sequence as shown in (2). They take values ± 1 with equal probability and are independent of each other by definition. Therefore, the probability mass function for x_{il}

for $l = 1, \dots, m_1$ can be expressed as

$$p_{x_{il}}(x) = \begin{cases} \gamma_i/2N_c^2 & x = \pm 1 \\ 1 - \gamma_i/N_c^2 & x = 0 \end{cases} \quad (12)$$

From (12), it is seen that $E\{x_{il}\} = 0$ and $\text{Var}\{x_{il}\} = \gamma_i/N_c^2$ for $l = 1, \dots, m_1$.

Although the stationary sequence $\{x_{il}\}_{l=1}^{m_1}$ is not independent, it is a 1-dependent random sequence⁴ and hence a ϕ -mixing random sequence satisfying the conditions in [4]. Therefore, central limit theorem (CLT) can be applied and for sufficiently large m_1 the probability distribution of x_i can be approximately expressed as $\mathcal{N}(0, m_1 \sigma^2)$ where

$$\sigma^2 = E\{x_{i1}^2\} + 2 \sum_{j=1}^{\infty} E\{x_{i1} x_{i(1+j)}\}. \quad (13)$$

It is easy to show that the cross-correlation term is zero. Therefore, the following asymptotic distribution is obtained:

$$x_i \sim \mathcal{N}(0, m_1 \gamma_i/N_c^2) \quad \gamma_i < N_c. \quad (14)$$

When $\gamma_i \geq N_c$, we can follow similar approaches and show that

$$x_i \sim \mathcal{N}(0, m_1/N_c) \quad \gamma_i \geq N_c. \quad (15)$$

For the MAI term, we can similarly show that for large correlation intervals, it can be approximated by the following Gaussian random variable:

$$a_i \sim \mathcal{N}(0, (N_u - 1)m_1/N_c). \quad (16)$$

B. Block Search Outputs

In this case, the received signal is correlated with shifted versions of the template signal (6). The block containing the signal cell is called the signal block while the blocks consisting of all non-signal cells are called non-signal blocks.

The correlation output can be expressed as

$$y_i = \int \tau(t) s_{m_2}^{(b)}(t - \Delta'_i) dt, \quad (17)$$

for $i = 1, \dots, B$ with $\Delta'_i = (i-1)KT_c$ and $N = KB$.

Again the output can be expressed as a sum of three terms, $y_i = x_i + a_i + n_i$, where x_i is the signal part or self interference term, a_i is MAI term and n_i is the output noise. Note that $n_i \sim \mathcal{N}(0, m_2 K \sigma_n^2)$. Also by similar CLT arguments, $a_i \sim \mathcal{N}(0, (N_u - 1)m_2 K/N_c)$.

For x_i , we have the following results:

1) *Non-Signal Block Case*: In this case, all of the K cells in the block are non-signal cells. The following lemma states that under some situations, the probability density function of the self interference terms of the non-signal block output can be approximated by a Gaussian density function.

Lemma 4.1: Assuming that each of the K cells in the block is away from the signal cell by more than N_c cells, the probability distribution of the self interference term of the non-signal block output can be approximated by the following Gaussian distribution for large values of correlation interval

⁴ (X_1, \dots, X_i) and (X_{i+j}, \dots) are independent whenever $j > 1$.

m_2

$$x_i \sim \mathcal{N}(0, m_2 K / N_c), \quad (18)$$

for $i = 1, \dots, l_b - 1, l_b + 1, \dots, B$, where l_b is the index of the signal block.

Proof See [2].

2) *Signal Block Case*: In this case, the signal cell is in the block and the remaining $K - 1$ cells are non-signal cells. For correlation over m_2 pulses, the distribution of the signal part of the signal block output can again be approximated by a Gaussian distribution under some conditions. We state the result for the simple case where $2 \leq K < N_c$ (see [2] for the general expression).

Lemma 4.2: For $2 \leq K < N_c$, the probability distribution of the signal part of the signal block output can be approximated by the following Gaussian distribution for large values of the correlation interval m_2

$$x_{l_b} \sim \mathcal{N} \left(m_2, m_2 \left[\frac{K^2 - 1}{3N_c^2} + \frac{n(n+1)(3K - 4n - 2)}{3KN_c^2} \right] \right), \quad (19)$$

where $n = \lfloor \frac{K-1}{2} \rfloor$. Or, in a more reduced form, $x_{l_b} \sim \mathcal{N}(m_2, m_2(5K^2 - c)/(12N_c^2))$ where $c = 5$ if K is odd and $c = 8$ for even K .

Proof See [2].

V. ACQUISITION TIME ANALYSIS

Using the approximate probability distributions derived in the previous section, the probabilities of false alarms and miss detections in different scenarios can be expressed using Q-functions. Then, the aim becomes the derivation of the mean acquisition time (MAT) formulas.

It is shown in [1] that when the signal cell location is uniformly distributed through the uncertainty region, the MAT for the serial search in a flat fading channel can be expressed approximately as

$$E\{T_{\text{serial}}\} \approx \frac{N(2 - P_D)(1 + CP_{FA}/m_1)}{2P_D} \cdot m_1 \quad (20)$$

for large N values, where N is the number of cells in the uncertainty region, P_D is the probability of detection of the signal path and P_{FA} is the probability of false alarm. The correlation over one frame is taken to be one unit time. When a false alarm occurs, C penalty time units elapse before the search continues with the next cell.

Using the signal flow graph approach in [1], we can show that for large number of blocks B , the MAT for the SBS algorithm can be expressed as

$$E\{T_{\text{SBS}}\} \approx \frac{[m_2 + KP_N(m_1 + CP_{FA})](2 - P_S P_D)B}{2P_S P_D}, \quad (21)$$

where P_N is the probability that a non-signal block output exceeds the threshold and P_S is the probability that a signal block output exceeds the threshold.

Define asymptotic gain as the ratio of the serial search MAT to the SBS MAT under the limiting conditions $\sigma_n^2 \rightarrow 0$ and $N_c \rightarrow \infty$. It can be shown [2] that

$$G = 10 \log \frac{N/2}{(B/2 + K/2)|_{B=K=\sqrt{N}}} = 10 \log \frac{\sqrt{N}}{2}. \quad (22)$$

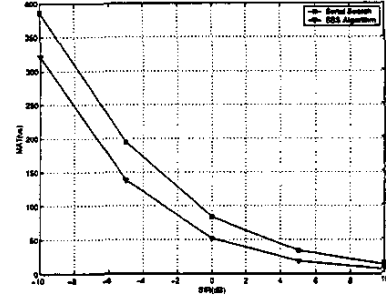


Fig. 1. Mean acquisition times for different algorithms.

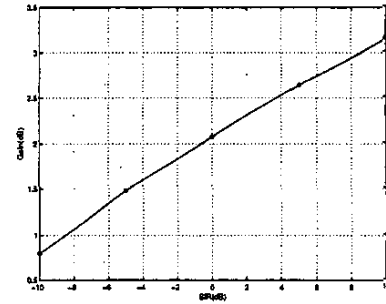


Fig. 2. Gain for different SIR values.

In other words, in the ideal conditions, the gain increases as the number of cells in the uncertainty region increases.

For realistic situations, we perform some simulation experiments to compare the algorithms. Figure 1 compares the MATs of serial search and SBS algorithms for different SIR⁵ values when the number of cells in the uncertainty region, N , is 360, chip interval T_c is 0.25ns, false alarm penalty is 100 unit time and there are 10 users in the system. Here, we can see that SBS algorithm has lower mean acquisition times and the gain increases as SIR increases as shown in Figure 2.

VI. CONCLUSION

In this paper, a rapid acquisition technique has been proposed and its performance has been evaluated. Its comparison to conventional serial search technique shows the efficiency of the algorithm.

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⁵For unit pulse energy, $\text{SIR} = -10 \log(\sigma_n^2 + (N_u - 1)/N_c)$ is used as the definition of the SIR.